# Package 'PAGFL'

February 20, 2025

**Title** Joint Estimation of Latent Groups and Group-Specific Coefficients in Panel Data Models

Version 1.1.3

Maintainer Paul Haimerl <paul.haimerl@econ.au.dk>

Description Latent group structures are a common challenge in panel data analysis. Disregarding group-level heterogeneity can introduce bias. Conversely, estimating individual coefficients for each cross-sectional unit is inefficient and may lead to high uncertainty. This package addresses the issue of unobservable group structures by implementing the pairwise adaptive group fused Lasso (PAGFL) by Mehrabani (2023) <doi:10.1016/j.jeconom.2022.12.002>. PAGFL identifies latent group structures and group-specific coefficients in a single step.

On top of that, we extend the PAGFL to time-varying coefficient functions.

```
License AGPL (>= 3)
Encoding UTF-8
RoxygenNote 7.3.2
```

LinkingTo Rcpp, RcppArmadillo, RcppParallel, RcppThread

Imports Rcpp, lifecycle, ggplot2, RcppParallel

BugReports https://github.com/Paul-Haimerl/PAGFL/issues

URL https://github.com/Paul-Haimerl/PAGFL

**Suggests** testthat (>= 3.0.0)

Config/testthat/edition 3

**NeedsCompilation** yes

Author Paul Haimerl [aut, cre] (<a href="https://orcid.org/0000-0003-3198-8317">https://orcid.org/0000-0003-3198-8317</a>), Stephan Smeekes [ctb] (<a href="https://orcid.org/0000-0002-0157-639X">https://orcid.org/0000-0003-3269-4601</a>), Ines Wilms [ctb] (<a href="https://orcid.org/0000-0002-1848-5582">https://orcid.org/0000-0002-1848-5582</a>)

**Depends** R (>= 3.5.0)

Repository CRAN

**Date/Publication** 2025-02-20 13:10:01 UTC

# **Contents**

grouped_plm				 															2
grouped_tv_plm .				 															6
pagfl																			9
sim_DGP																			14
sim_tv_DGP																			16
tv_pagfl																			19

24

grouped\_plm

Grouped Panel Data Model

## Description

**Index** 

Estimate a grouped panel data model given an observed group structure. Slope parameters are homogeneous within groups but heterogeneous across groups. This function supports both static and dynamic panel data models, with or without endogenous regressors.

## Usage

```
grouped_plm(
  formula,
  data,
  groups,
  index = NULL,
  n_periods = NULL,
 method = "PLS",
  Z = NULL,
 bias_correc = FALSE,
  rho = 0.07 * log(N * n_periods)/sqrt(N * n_periods),
  verbose = TRUE,
  parallel = TRUE,
)
## S3 method for class 'gplm'
print(x, ...)
## S3 method for class 'gplm'
formula(x, ...)
## S3 method for class 'gplm'
df.residual(object, ...)
## S3 method for class 'gplm'
summary(object, ...)
```

```
## S3 method for class 'gplm'
coef(object, ...)
## S3 method for class 'gplm'
residuals(object, ...)
## S3 method for class 'gplm'
fitted(object, ...)
```

#### **Arguments**

formula a formula object describing the model to be estimated.

data a data.frame or matrix holding a panel data set. If no index variables are

provided, the panel must be balanced and ordered in the long format Y = $(Y'_1, \ldots, Y'_N)', Y_i = (Y_{i1}, \ldots, Y_{iT})'$  with  $Y_{it} = (y_{it}, x'_{it})'$ . Conversely, if data is not ordered or not balanced, data must include two index variables that de-

clare the cross-sectional unit i and the time period t of each observation.

a numerical or character vector of length N that indicates the group membership groups

of each cross-sectional unit i.

index a character vector holding two strings. The first string denotes the name of

the index variable identifying the cross-sectional unit i and the second string represents the name of the variable declaring the time period t. The data is automatically sorted according to the variables in index, which may produce errors when the time index is a character variable. In case of a balanced panel data set that is ordered in the long format, index can be left empty if the the

number of time periods n\_periods is supplied.

the number of observed time periods T. If an index is passed, this argument n\_periods

can be left empty.

method the estimation method. Options are

> "PLS" for using the penalized least squares (PLS) algorithm. We recommend PLS in case of (weakly) exogenous regressors (Mehrabani, 2023, sec. 2.2).

> "PGMM" for using the penalized Generalized Method of Moments (PGMM). PGMM is required when instrumenting endogenous regressors, in which case a matrix Z containing the necessary exogenous instruments must be supplied (Mehrabani, 2023, sec. 2.3).

Default is "PLS".

a  $NT \times q$  matrix or data.frame of exogenous instruments, where  $q \geq p$ ,  $Z = (z_1, \ldots, z_N)', z_i = (z_{i1}, \ldots, z_{iT})'$  and  $z_{it}$  is a  $q \times 1$  vector. Z is only required when method = "PGMM" is selected. When using "PLS", the argument

can be left empty or it is disregarded. Default is NULL.

logical. If TRUE, a Split-panel Jackknife bias correction following Dhaene and Jochmans (2015) is applied to the slope parameters. We recommend using the

correction when working with dynamic panels. Default is FALSE.

Ζ

bias\_correc

rho a tuning parameter balancing the fitness and penalty terms in the IC. If left unspecified, the heuristic  $\rho=0.07\frac{\log(NT)}{\sqrt{NT}}$  of Mehrabani (2023, sec. 6) is used. We recommend the default.

verbose logical. If TRUE, helpful warning messages are shown. Default is TRUE.

parallel logical. If TRUE, certain operations are parallelized across multiple cores. De-

fault is TRUE.

... ellipsis

x of class gplm.

object of class gplm.

## **Details**

Consider the grouped panel data model

$$y_{it} = \gamma_i + \beta_i' x_{it} + \epsilon_{it}, \quad i = 1, \dots, N, \ t = 1, \dots, T,$$

where  $y_{it}$  is the scalar dependent variable,  $\gamma_i$  is an individual fixed effect,  $x_{it}$  is a  $p \times 1$  vector of explanatory variables, and  $\epsilon_{it}$  is a zero mean error. The coefficient vector  $\beta_i$  is subject to the observed group pattern

$$\beta_i = \sum_{k=1}^K \alpha_k \mathbf{1} \{ i \in G_k \},$$

with  $\bigcup_{k=1}^K G_k = \{1,\ldots,N\}$ ,  $G_k \cap G_j = \emptyset$  and  $\|\alpha_k - \alpha_j\| \neq 0$  for any  $k \neq j, k = 1,\ldots,K$ .

Using PLS, the group-specific coefficients for group k are obtained via OLS

$$\hat{\alpha}_k = \left(\sum_{i \in G_k} \sum_{t=1}^T \tilde{x}_{it} \tilde{x}'_{it}\right)^{-1} \sum_{i \in G_k} \sum_{t=1}^T \tilde{x}_{it} \tilde{y}_{it},$$

where  $\tilde{a}_{it} = a_{it} - T^{-1} \sum_{t=1}^{T} a_{it}$ ,  $a = \{y, x\}$  to concentrate out the individual fixed effects  $\gamma_i$  (within-transformation).

In case of PGMM, the slope coefficients are derived as

$$\hat{\alpha}_k = \left( \left[ \sum_{i \in G_k} T^{-1} \sum_{t=1}^T z_{it} \Delta x_{it} \right]' W_k \left[ \sum_{i \in G_k} T^{-1} \sum_{t=1}^T z_{it} \Delta x_{it} \right] \right)^{-1}$$

$$\left[\sum_{i \in G_k} T^{-1} \sum_{t=1}^T z_{it} \Delta x_{it}\right]' W_k \left[\sum_{i \in G_k} T^{-1} \sum_{t=1}^T z_{it} \Delta y_{it}\right],$$

where  $W_k$  is a  $q \times q$  p.d. symmetric weight matrix and  $\Delta$  denotes the first difference operator  $\Delta x_{it} = x_{it} - x_{it-1}$  (first-difference transformation).

#### Value

An object of class gplm holding

model a data.frame containing the dependent and explanatory variables as well as

cross-sectional and time indices,

coefficients a  $K \times p$  matrix of the group-specific parameter estimates,

groups a list containing (i) the total number of groups K and (ii) a vector of group

memberships  $g_1, \ldots, g_N$ ), where  $g_i = k$  if i is assigned to group k,

residuals a vector of residuals of the demeaned model,

fitted a vector of fitted values of the demeaned model,

args a list of additional arguments,

IC a list containing (i) the value of the IC and (ii) the MSE,

call the function call.

A gplm object has print, summary, fitted, residuals, formula, df.residual, and coef S3 methods.

## Author(s)

Paul Haimerl

#### References

Dhaene, G., & Jochmans, K. (2015). Split-panel jackknife estimation of fixed-effect models. *The Review of Economic Studies*, 82(3), 991-1030. doi:10.1093/restud/rdv007. Mehrabani, A. (2023). Estimation and identification of latent group structures in panel data. *Journal of Econometrics*, 235(2), 1464-1482. doi:10.1016/j.jeconom.2022.12.002.

## Examples

```
# Simulate a panel with a group structure
set.seed(1)
sim < -sim_DGP(N = 20, n_periods = 80, p = 2, n_groups = 3)
y <- sim$y
X <- sim$X
groups <- sim$groups
df \leftarrow cbind(y = c(y), X)
# Estimate the grouped panel data model
estim <- grouped_plm(y ~ ., data = df, groups = groups, n_periods = 80, method = "PLS")
summary(estim)
# Lets pass a panel data set with explicit cross-sectional and time indicators
i_index <- rep(1:20, each = 80)
t_{index} < - rep(1:80, 20)
df <- data.frame(y = c(y), X, i_index = i_index, t_index = t_index)</pre>
estim <- grouped_plm(</pre>
 data = df, index = c("i_index", "t_index"), groups = groups, method = "PLS"
```

grouped\_tv\_plm

```
)
summary(estim)
```

grouped\_tv\_plm

Grouped Time-varying Panel Data Model

## Description

Estimate a grouped time-varying panel data model given an observed group structure. Coefficient functions are homogeneous within groups but heterogeneous across groups. The time-varying coefficients are modeled as polynomial B-splines. The function supports both static and dynamic panel data models.

#### Usage

```
grouped_tv_plm(
  formula,
  data,
  groups,
  index = NULL,
  n_periods = NULL,
  d = 3,
 M = floor(length(y)^(1/7) - log(p)),
  const_coef = NULL,
  rho = 0.04 * log(N * n_periods)/sqrt(N * n_periods),
  verbose = TRUE,
  parallel = TRUE,
)
## S3 method for class 'tv_gplm'
summary(object, ...)
## S3 method for class 'tv_gplm'
formula(x, ...)
## S3 method for class 'tv_gplm'
df.residual(object, ...)
## S3 method for class 'tv_gplm'
print(x, ...)
## S3 method for class 'tv_gplm'
coef(object, ...)
## S3 method for class 'tv_gplm'
residuals(object, ...)
```

grouped\_tv\_plm 7

```
## S3 method for class 'tv_gplm'
fitted(object, ...)
```

#### **Arguments**

formula a formula object describing the model to be estimated.

data a data. frame or matrix holding a panel data set. If no index variables are

provided, the panel must be balanced and ordered in the long format  $Y = (Y_1', \ldots, Y_N')'$ ,  $Y_i = (Y_{i1}, \ldots, Y_{iT})'$  with  $Y_{it} = (y_{it}, x_{it}')'$ . Conversely, if data is not ordered or not balanced, data must include two index variables that de-

clare the cross-sectional unit i and the time period t of each observation.

groups a numerical or character vector of length N that indicates the group membership

of each cross-sectional unit i.

index a character vector holding two strings. The first string denotes the name of

the index variable identifying the cross-sectional unit i, and the second string represents the name of the variable declaring the time period t. The data is automatically sorted according to the variables in index, which may produce errors when the time index is a character variable. In case of a balanced panel data set that is ordered in the long format, index can be left empty if the the

number of time periods n\_periods is supplied.

n\_periods the number of observed time periods T. If an index character vector is passed,

this argument can be left empty. Default is Null.

d the polynomial degree of the B-splines. Default is 3.

M the number of interior knots of the B-splines. If left unspecified, the default

heuristic  $M = \text{floor}((NT)^{\frac{1}{7}} - \log(p))$  is used. Note that M does not include the boundary knots and the entire sequence of knots is of length M + d + 1.

const\_coef a character vector containing the variable names of explanatory variables that

enter with time-constant coefficients.

rho the tuning parameter balancing the fitness and penalty terms in the IC. If left

unspecified, the heuristic  $\rho = 0.07 \frac{\log(NT)}{\sqrt{NT}}$  of Mehrabani (2023, sec. 6) is used.

We recommend the default.

verbose logical. If TRUE, helpful warning messages are shown. Default is TRUE.

parallel logical. If TRUE, certain operations are parallelized across multiple cores. De-

fault is TRUE.

... ellipsis

object of class tv\_gplm.

x of class tv\_gplm.

#### **Details**

Consider the grouped time-varying panel data model

$$y_{it} = \gamma_i + \beta'_i(t/T)x_{it} + \epsilon_{it}, \quad i = 1, ..., N, \ t = 1, ..., T,$$

8 grouped\_tv\_plm

where  $y_{it}$  is the scalar dependent variable,  $\gamma_i$  is an individual fixed effect,  $x_{it}$  is a  $p \times 1$  vector of explanatory variables, and  $\epsilon_{it}$  is a zero mean error. The coefficient vector  $\beta_i(t/T)$  is subject to the observed group pattern

$$\beta_i\left(\frac{t}{T}\right) = \sum_{k=1}^K \alpha_k\left(\frac{t}{T}\right) \mathbf{1}\{i \in G_k\},\,$$

with  $\bigcup_{k=1}^K G_k = \{1, \dots, N\}$ ,  $G_k \cap G_j = \emptyset$  and  $\|\alpha_k - \alpha_j\| \neq 0$  for any  $k \neq j, k = 1, \dots, K$ .

 $\alpha_k(t/T)$  and, in turn,  $\beta_i(t/T)$  is estimated as polynomial B-splines using the penalized sieve-technique. To this end, let B(v) denote a M+d+1 vector of polynomial spline basis functions, where d represents the polynomial degree and M gives the number of interior knots of the B-spline.  $\alpha_k(t/T)$  is approximated by forming a linear combination of the basis functions  $\alpha_k(t/T) \approx \xi_k' B(t/T)$ , where  $\xi_k$  is a  $(M+d+1) \times p$  coefficient matrix.

The explanatory variables are projected onto the spline basis system, which results in the  $(M+d+1)p \times 1$  vector  $z_{it} = x_{it} \otimes B(v)$ . Subsequently, the DGP can be reformulated as

$$y_{it} = \gamma_i + z'_{it} \operatorname{vec}(\pi_i) + u_{it},$$

where  $\pi_i = \xi_k$  if  $i \in G_k$ ,  $u_{it} = \epsilon_{it} + \eta_{it}$ , and  $\eta_{it}$  reflects a sieve approximation error. We refer to Su et al. (2019, sec. 2) for more details on the sieve technique.

Finally,  $\hat{\alpha}_k(t/T)$  is obtained as  $\hat{\alpha}_k(t/T) = \hat{\xi}_k' B(t/T)$ , where the vector of control points  $\xi_k$  is estimated using *OLS* 

$$\hat{\xi}_k = \left(\sum_{i \in G_k} \sum_{t=1}^T \tilde{z}_{it} \tilde{z}'_{it}\right)^{-1} \sum_{i \in G_k} \sum_{t=1}^T \tilde{z}_{it} \tilde{y}_{it},$$

and  $\tilde{a}_{it} = a_{it} - T^{-1} \sum_{t=1}^{T} a_{it}$ ,  $a = \{y, z\}$  to concentrate out the fixed effect  $\gamma_i$  (within-transformation).

In case of an unbalanced panel data set, the earliest and latest available observations per group define the start and end-points of the interval on which the group-specific time-varying coefficients are defined.

## Value

An object of class tv\_gplm holding

model a data.frame containing the dependent and explanatory variables as well as

cross-sectional and time indices,

coefficients  $\det p^{(1)}$  denote the number of time-varying and  $p^{(2)}$  the number of time constant

coefficients. A list holding (i) a  $T \times p^{(1)} \times K$  array of the group-specific functional coefficients and (ii) a  $K \times p^{(2)}$  matrix of time-constant estimates.

groups a list containing (i) the total number of groups K and (ii) a vector of group

memberships  $(\hat{g}_1, \dots, \hat{g}_N)$ , where  $\hat{g}_i = k$  if i is part of group k,

residuals a vector of residuals of the demeaned model,

fitted a vector of fitted values of the demeaned model,

args a list of additional arguments,

IC a list containing (i) the value of the IC and (ii) the MSE,

call the function call.

An object of class tv\_gplm has print, summary, fitted, residuals, formula, df.residual and coef S3 methods.

## Author(s)

Paul Haimerl

#### References

Su, L., Wang, X., & Jin, S. (2019). Sieve estimation of time-varying panel data models with latent structures. *Journal of Business & Economic Statistics*, 37(2), 334-349. doi:10.1080/07350015.2017.1340299.

## **Examples**

```
# Simulate a time-varying panel with a trend and a group pattern
set.seed(1)
sim <- sim_tv_DGP(N = 10, n_periods = 50, intercept = TRUE, p = 2)
df <- data.frame(y = c(sim$y), X = sim$X)
groups <- sim$groups
# Estimate the time-varying grouped panel data model
estim <- grouped_tv_plm(y ~ ., data = df, n_periods = 50, groups = groups)
summary(estim)</pre>
```

pagfl

Pairwise Adaptive Group Fused Lasso

## Description

Estimate panel data models with a latent group structure using the pairwise adaptive group fused Lasso (*PAGFL*) by Mehrabani (2023). The *PAGFL* jointly identifies the group structure and group-specific slope parameters. The function supports both static and dynamic panels, with or without endogenous regressors.

## Usage

```
pagfl(
  formula,
  data,
  index = NULL,
  n_periods = NULL,
  lambda,
  method = "PLS",
  Z = NULL,
  min_group_frac = 0.05,
  bias_correc = FALSE,
  kappa = 2,
  max_iter = 5000,
  tol_convergence = 1e-08,
  tol_group = 0.001,
  rho = 0.07 * log(N * n_periods)/sqrt(N * n_periods),
```

```
varrho = max(sqrt(5 * N * n_periods * p)/log(N * n_periods * p) - 7, 1),
  verbose = TRUE,
  parallel = TRUE,
)
## S3 method for class 'pagfl'
print(x, ...)
## S3 method for class 'pagfl'
formula(x, ...)
## S3 method for class 'pagfl'
df.residual(object, ...)
## S3 method for class 'pagfl'
summary(object, ...)
## S3 method for class 'pagfl'
coef(object, ...)
## S3 method for class 'pagfl'
residuals(object, ...)
## S3 method for class 'pagfl'
fitted(object, ...)
```

## **Arguments**

formula a formula object describing the model to be estimated.

data a data.frame or matrix holding a panel data set. If no index variables are

provided, the panel must be balanced and ordered in the long format  $Y = (Y_1', \ldots, Y_N')'$ ,  $Y_i = (Y_{i1}, \ldots, Y_{iT})'$  with  $Y_{it} = (y_{it}, x_{it}')'$ . Conversely, if data is not ordered or not balanced, data must include two index variables that de-

clare the cross-sectional unit i and the time period t of each observation.

index a character vector holding two strings. The first string denotes the name of

the index variable identifying the cross-sectional unit i and the second string represents the name of the variable declaring the time period t. The data is automatically sorted according to the variables in index, which may produce errors when the time index is a character variable. In case of a balanced panel data set that is ordered in the long format, index can be left empty if the the

number of time periods n\_periods is supplied.

 $n_periods$  the number of observed time periods T. If an index character vector is passed,

this argument can be left empty.

lambda the tuning parameter determining the strength of the penalty term. Either a single

 $\lambda$  or a vector of candidate values can be passed. If a vector is supplied, a BIC-

type IC automatically selects the best fitting  $\lambda$  value.

method the estimation method. Options are

"PLS" for using the penalized least squares (*PLS*) algorithm. We recommend *PLS* in case of (weakly) exogenous regressors (Mehrabani, 2023, sec. 2.2).

"PGMM" for using the penalized Generalized Method of Moments (*PGMM*). *PGMM* is required when instrumenting endogenous regressors, in which case a matrix **Z** containing the necessary exogenous instruments must be supplied (Mehrabani, 2023, sec. 2.3).

Default is "PLS".

a  $NT \times q$  matrix or data.frame of exogenous instruments, where  $q \geq p$ ,  $\mathbf{Z} = (z_1, \ldots, z_N)'$ ,  $z_i = (z_{i1}, \ldots, z_{iT})'$  and  $z_{it}$  is a  $q \times 1$  vector. Z is only required when method = "PGMM" is selected. When using "PLS", either pass NULL or Z is disregarded. Default is NULL.

min\_group\_frac the minimum group cardinality as a fraction of the total number of individuals N. In case a group falls short of this threshold, each of its members is allocated to one of the remaining groups according to the MSE. Default is 0.05.

bias\_correc logical. If TRUE, a Split-panel Jackknife bias correction following Dhaene and Jochmans (2015) is applied to the slope parameters. We recommend using the correction when working with dynamic panels. Default is FALSE.

kappa the a non-negative weight used to obtain the adaptive penalty weights. Default is 2.

the maximum number of iterations for the *ADMM* estimation algorithm. Default is  $1 * 10^4$ .

tol\_convergence

max\_iter

tol\_group

rho

varrho

Ζ

the tolerance limit for the stopping criterion of the iterative *ADMM* estimation algorithm. Default is  $1*10^{-8}$ .

the tolerance limit for within-group differences. Two individuals i, j are assigned to the same group if the Frobenius norm of their coefficient vector difference is below this threshold. Default is  $1 * 10^{-3}$ .

the tuning parameter balancing the fitness and penalty terms in the IC that determines the penalty parameter  $\lambda$ . If left unspecified, the heuristic  $\rho=0.07\frac{\log(NT)}{\sqrt{NT}}$  of Mehrabani (2023, sec. 6) is used. We recommend the default.

the non-negative Lagrangian *ADMM* penalty parameter. For *PLS*, the  $\varrho$  value is trivial. However, for *PGMM*, small values lead to slow convergence. If left unspecified, the default heuristic  $\varrho = \max(\frac{\sqrt{5NTp}}{\log(NTp)} - 7, 1)$  is used.

verbose logical. If TRUE, helpful warning messages are shown. Default is TRUE.

parallel logical. If TRUE, certain operations are parallelized across multiple cores. Default is TRUE.

... ellipsis

x of class pagfl.
object of class pagfl.

#### **Details**

Consider the grouped panel data model

$$y_{it} = \gamma_i + \beta_i' x_{it} + \epsilon_{it}, \quad i = 1, \dots, N, \ t = 1, \dots, T,$$

where  $y_{it}$  is the scalar dependent variable,  $\gamma_i$  is an individual fixed effect,  $x_{it}$  is a  $p \times 1$  vector of weakly exogenous explanatory variables, and  $\epsilon_{it}$  is a zero mean error. The coefficient vector  $\beta_i$  is subject to the latent group pattern

$$\beta_i = \sum_{k=1}^K \alpha_k \mathbf{1} \{ i \in G_k \},$$

with  $\bigcup_{k=1}^K G_k = \{1, \dots, N\}$ ,  $G_k \cap G_j = \emptyset$  and  $\|\alpha_k - \alpha_j\| \neq 0$  for any  $k \neq j, k = 1, \dots, K$ .

The *PLS* method jointly estimates the latent group structure and group-specific coefficients by minimizing the criterion

$$Q_{NT}(\boldsymbol{\beta}, \lambda) = \frac{1}{T} \sum_{i=1}^{N} \sum_{t=1}^{T} (\tilde{y}_{it} - \beta_i' \tilde{x}_{it})^2 + \frac{\lambda}{N} \sum_{i=1}^{N-1} \sum_{j>i}^{N} \dot{\omega}_{ij} \|\beta_i - \beta_j\|$$

with respect to  $\beta = (\beta_1', \dots, \beta_N')'$ .  $\tilde{a}_{it} = a_{it} - T^{-1} \sum_{t=1}^T a_{it}$ ,  $a = \{y, x\}$  to concentrate out the individual fixed effects  $\gamma_i$ .  $\lambda$  is the penalty tuning parameter and  $\dot{\omega}_{ij}$  reflects adaptive penalty weights (see Mehrabani, 2023, eq. 2.6).  $\|\cdot\|$  denotes the Frobenius norm. The adaptive weights  $\dot{w}_{ij}$  are obtained by a preliminary individual least squares estimation. The criterion function is minimized via an iterative alternating direction method of multipliers (*ADMM*) algorithm (see Mehrabani, 2023, sec. 5.1).

*PGMM* employs a set of instruments Z to control for endogenous regressors. Using *PGMM*,  $\beta$  is estimated by minimizing

$$Q_{NT}(\boldsymbol{\beta}, \lambda) = \sum_{i=1}^{N} \left[ \frac{1}{N} \sum_{t=1}^{T} z_{it} (\Delta y_{it} - \beta_i' \Delta x_{it}) \right]' W_i \left[ \frac{1}{T} \sum_{t=1}^{T} z_{it} (\Delta y_{it} - \beta_i' \Delta x_{it}) \right] + \frac{\lambda}{N} \sum_{i=1}^{N-1} \sum_{j>i}^{N} \ddot{\omega}_{ij} \|\beta_i - \beta_j\|.$$

 $\ddot{\omega}_{ij}$  are obtained by an initial *GMM* estimation.  $\Delta$  gives the first differences operator  $\Delta y_{it} = y_{it} - y_{it-1}$ .  $W_i$  represents a data-driven  $q \times q$  weight matrix. I refer to Mehrabani (2023, eq. 2.10) for more details. Again, the criterion function is minimized using an efficient *ADMM* algorithm (Mehrabani, 2023, sec. 5.2).

Two individuals are assigned to the same group if  $\|\hat{\beta}_i - \hat{\beta}_j\| \le \epsilon_{\text{tol}}$ , where  $\epsilon_{\text{tol}}$  is determined by tol\_group. Subsequently, the number of groups follows as the number of distinct elements in  $\hat{\beta}$ . Given an estimated group structure, it is straightforward to obtain post-Lasso estimates using group-wise least squares or *GMM* (see grouped\_plm).

We recommend identifying a suitable  $\lambda$  parameter by passing a logarithmically spaced grid of candidate values with a lower limit close to 0 and an upper limit that leads to a fully homogeneous panel. A BIC-type information criterion then selects the best fitting  $\lambda$  value.

#### Value

An object of class pagf1 holding

model a data. frame containing the dependent and explanatory variables as well as

cross-sectional and time indices,

coefficients a  $\hat{K} \times p$  matrix of the post-Lasso group-specific parameter estimates,

groups a list containing (i) the total number of groups  $\hat{K}$  and (ii) a vector of estimated

group memberships  $(\hat{g}_1, \dots, \hat{g}_N)$ , where  $\hat{g}_i = k$  if i is assigned to group k,

residuals a vector of residuals of the demeaned model, fitted a vector of fitted values of the demeaned model,

args a list of additional arguments,

IC a list containing (i) the value of the IC, (ii) the employed tuning parameter  $\lambda$ ,

and (iii) the MSE,

convergence a list containing (i) a logical variable indicating if convergence was achieved

and (ii) the number of executed ADMM algorithm iterations,

call the function call.

A pagfl object has print, summary, fitted, residuals, formula, df.residual, and coef S3 methods.

## Author(s)

Paul Haimerl

#### References

Dhaene, G., & Jochmans, K. (2015). Split-panel jackknife estimation of fixed-effect models. *The Review of Economic Studies*, 82(3), 991-1030. doi:10.1093/restud/rdv007. Mehrabani, A. (2023). Estimation and identification of latent group structures in panel data. *Journal of Econometrics*, 235(2), 1464-1482. doi:10.1016/j.jeconom.2022.12.002.

## **Examples**

```
# Simulate a panel with a group structure
set.seed(1)
sim <- sim_DGP(N = 20, n_periods = 80, p = 2, n_groups = 3)
y <- sim$y
X <- sim$X
df <- cbind(y = c(y), X)

# Run the PAGFL procedure
estim <- pagfl(y ~ ., data = df, n_periods = 80, lambda = 0.5, method = "PLS")
summary(estim)

# Lets pass a panel data set with explicit cross-sectional and time indicators
i_index <- rep(1:20, each = 80)
t_index <- rep(1:80, 20)
df <- data.frame(y = c(y), X, i_index = i_index, t_index = t_index)</pre>
```

14 sim\_DGP

```
estim <- pagfl(
   y ~ .,
   data = df, index = c("i_index", "t_index"), lambda = 0.5, method = "PLS"
)
summary(estim)</pre>
```

sim\_DGP

Simulate a Panel With a Group Structure in the Slope Coefficients

## Description

Construct a static or dynamic, exogenous or endogenous panel data set subject to a group structure in the slope coefficients with optional AR(1) or GARCH(1,1) innovations.

## Usage

```
sim_DGP(
  N = 50,
  n_periods = 40,
  p = 2,
  n_groups = 3,
  group_proportions = NULL,
  error_spec = "iid",
  dynamic = FALSE,
  dyn_panel = lifecycle::deprecated(),
  q = NULL,
  alpha_0 = NULL
)
```

#### **Arguments**

```
the number of cross-sectional units. Default is 50.
n_periods
                  the number of simulated time periods T. Default is 40.
                  the number of explanatory variables. Default is 2.
n_groups
                  the number of groups K. Default is 3.
group_proportions
                  a numeric vector of length n_groups indicating size of each group as a fraction
                  of N. If NULL, all groups are of size N/K. Default is NULL.
                   options include
error_spec
                   "iid" for iid errors.
                   "AR" for an AR(1) error process with an autoregressive coefficient of 0.5.
                   "GARCH" for a GARCH(1, 1) error process with a 0.05 constant, a 0.05 ARCH
                       and a 0.9 GARCH coefficient.
                  Default is "iid".
```

sim\_DGP 15

dynamic Logical. If TRUE, the panel includes one stationary autoregressive lag of  $y_{it}$  as an explanatory variable (see sec. Details for more information on the AR coefficient). Default is FALSE.

dyn\_panel [Deprecated] deprecated and replaced by dynamic.

q the number of exogenous instruments when a panel with endogenous regressors is to be simulated. If panel data set with exogenous regressors is supposed to be generated, pass NULL. Default is NULL.

alpha\_0 a  $K \times p$  matrix of group-specific coefficients. If dynamic = TRUE, the first column represents the stationary AR coefficient. If NULL, the coefficients are drawn randomly (see sec. Details). Default is NULL.

#### **Details**

The scalar dependent variable  $y_{it}$  is generated according to the following grouped panel data model

$$y_{it} = \gamma_i + \beta_i' x_{it} + u_{it}, \quad i = \{1, \dots, N\}, \quad t = \{1, \dots, T\}.$$

 $\gamma_i$  represents individual fixed effects and  $x_{it}$  a  $p \times 1$  vector of regressors. The individual slope coefficient vectors  $\beta_i$  are subject to a group structure

$$\beta_i = \sum_{k=1}^K \alpha_k \mathbf{1} \{ i \in G_k \},$$

with  $\bigcup_{k=1}^K G_k = \{1, \dots, N\}$ ,  $G_k \cap G_j = \emptyset$  and  $\|\alpha_k - \alpha_j\| \neq 0$  for any  $k \neq j$ ,  $k = 1, \dots, K$ . The total number of groups K is determined by n\_groups.

If a panel data set with exogenous regressors is generated (set q = NULL), the explanatory variables are simulated according to

$$x_{it,j} = 0.2\gamma_i + e_{it,j}, \quad \gamma_i, e_{it,j} \sim i.i.d.N(0,1), \quad j = \{1, \dots, p\},\$$

where  $e_{it,j}$  denotes a series of innovations.  $\gamma_i$  and  $e_i$  are independent of each other.

In case alpha\_0 = NULL, the group-level slope parameters  $\alpha_k$  are drawn from  $\sim U[-2,2]$ .

If a dynamic panel is specified (dynamic = TRUE), the AR coefficients  $\beta_i^{AR}$  are drawn from a uniform distribution with support (-1,1) and  $x_{it,j}=e_{it,j}$ . Moreover, the individual fixed effects enter the dependent variable via  $(1-\beta_i^{AR})\gamma_i$  to account for the autoregressive dependency. We refer to Mehrabani (2023, sec 6) for details.

When specifying an endogenous panel (set q to  $q \ge p$ ), the  $e_{it,j}$  correlate with the cross-sectional innovations  $u_{it}$  by a magnitude of 0.5 to produce endogenous regressors ( $\mathrm{E}(u|X) \ne 0$ ). However, the endogenous regressors can be accounted for by exploiting the q instruments in Z, for which  $\mathrm{E}(u|Z) = 0$  holds. The instruments and the first stage coefficients are generated in the same fashion as X and  $\alpha$  when  $q = \mathrm{NULL}$ .

The function nests, among other, the DGPs employed in the simulation study of Mehrabani (2023, sec. 6).

sim\_tv\_DGP

## Value

Α	list	ho	lding	3

alpha	the $K\times p$ matrix of group-specific slope parameters. If dynamic = TRUE, the first column holds the $AR$ coefficient.
groups	a vector indicating the group memberships $(g_1,\ldots,g_N)$ , where $g_i=k$ if $i\in \operatorname{group} k$ .
у	a $NT \times 1$ vector of the dependent variable, with $\mathbf{y} = (y_1, \dots, y_N)', y_i = (y_{i1}, \dots, y_{iT})'$ and the scalar $y_{it}$ .
Χ	a $NT \times p$ matrix of explanatory variables, with $\boldsymbol{X} = (x_1, \dots, x_N)', x_i = (x_{i1}, \dots, x_{iT})'$ and the $p \times 1$ vector $x_{it}$ .
Z	a $NT \times q$ matrix of instruments , where $q \geq p$ , $\boldsymbol{Z} = (z_1, \ldots, z_N)'$ , $z_i = (z_{i1}, \ldots, z_{iT})'$ and $z_{it}$ is a $q \times 1$ vector. In case a panel with exogenous regressors is generated (q = NULL), $\boldsymbol{Z}$ equals NULL.
data	a $NT \times (p+1)$ data.frame of the outcome and the explanatory variables.

## Author(s)

Paul Haimerl

## References

Mehrabani, A. (2023). Estimation and identification of latent group structures in panel data. *Journal of Econometrics*, 235(2), 1464-1482. doi:10.1016/j.jeconom.2022.12.002.

## **Examples**

```
# Simulate DGP 1 from Mehrabani (2023, sec. 6)
set.seed(1)
alpha_0_DGP1 <- matrix(c(0.4, 1, 1.6, 1.6, 1, 0.4), ncol = 2)
DGP1 <- sim_DGP(
    N = 50, n_periods = 20, p = 2, n_groups = 3,
    group_proportions = c(.4, .3, .3), alpha_0 = alpha_0_DGP1
)</pre>
```

 $sim_tv_DGP$ 

Simulate a Time-varying Panel With a Group Structure in the Slope Coefficients

## Description

Construct a time-varying panel data set subject to a group structure in the slope coefficients with optional AR(1) innovations.

sim\_tv\_DGP

## Usage

```
sim_tv_DGP(
  N = 50,
  n_periods = 40,
  intercept = TRUE,
  p = 1,
  n_groups = 3,
  d = 3,
  dynamic = FALSE,
  group_proportions = NULL,
  error_spec = "iid",
  locations = NULL,
  scales = NULL,
  polynomial_coef = NULL,
  sd_error = 1
)
```

## **Arguments**

N the number of cross-sectional units. Default is 50.

n\_periods the number of simulated time periods T. Default is 40.

intercept logical. If TRUE, a time-varying intercept is generated.

p the number of simulated explanatory variables

n\_groups the number of groups K. Default is 3.

d the polynomial degree used to construct the time-varying coefficients.

dynamic Logical. If TRUE, the panel includes one stationary autoregressive lag of  $y_{it}$  as a

regressor. Default is FALSE.

group\_proportions

a numeric vector of length n\_groups indicating size of each group as a fraction

of N. If NULL, all groups are of size N/K. Default is NULL.

error\_spec options include

"iid" for iid errors.

"AR" for an AR(1) error process with an autoregressive coefficient of 0.5.

Default is "iid".

locations a  $p \times K$  matrix of location parameters of a logistic distribution function used

to construct the time-varying coefficients. If left empty, the location parameters

are drawn randomly. Default is NULL.

scales a  $p \times K$  matrix of scale parameters of a logistic distribution function used to

construct the time-varying coefficients. If left empty, the location parameters

are drawn randomly. Default is NULL.

polynomial\_coef

a  $p \times d \times K$  array of coefficients for a the polynomials used to construct the timevarying coefficients. If left empty, the location parameters are drawn randomly.

Default is NULL.

sd\_error standard deviation of the cross-sectional errors. Default is 1.

18 sim\_tv\_DGP

#### **Details**

The scalar dependent variable  $y_{it}$  is generated according to the following time-varying grouped panel data model

$$y_{it} = \gamma_i + \beta'_{it} x_{it} + u_{it}, \quad i = 1, \dots, N, \ t = 1, \dots, T,$$

where  $\gamma_i$  is an individual fixed effect and  $x_{it}$  is a  $p \times 1$  vector of explanatory variables. The coefficient vector  $\beta_i = \{\beta'_{i1}, \dots, \beta'_{iT}\}'$  is subject to the group pattern

$$\beta_i\left(\frac{t}{T}\right) = \sum_{k=1}^K \alpha_k\left(\frac{t}{T}\right) \mathbf{1}\{i \in G_k\},$$

with  $\bigcup_{k=1}^K G_k = \{1,\ldots,N\}$ ,  $G_k \cap G_j = \emptyset$  and  $\sup_{v \in [0,1]} (\|\alpha_k(v) - \alpha_j(v)\|) \neq 0$  for any  $k \neq j$ ,  $k = 1,\ldots,K$ . The total number of groups K is determined by n\_groups.

The predictors are simulated as:

$$x_{it,j} = 0.2\gamma_i + e_{it,j}, \quad \gamma_i, e_{it,j} \sim i.i.d.N(0,1), \quad j = \{1, \dots, p\},\$$

where  $e_{it,j}$  denotes a series of innovations.  $\gamma_i$  and  $e_i$  are independent of each other.

The errors  $u_{it}$  feature a iid standard normal distribution.

In case locations = NULL, the location parameters are drawn from  $\sim U[0.3,0.9]$ . In case scales = NULL, the scale parameters are drawn from  $\sim U[0.01,0.09]$ . In case polynomial\_coef = NULL, the polynomial coefficients are drawn from  $\sim U[-20,20]$  and normalized so that all coefficients of one polynomial sum up to 1. The final coefficient function follows as  $\alpha_k(t/T) = 3*F(t/T,location,scale) + \sum_{j=1}^d a_j(t/T)^j$ , where  $F(\cdot,location,scale)$  denotes a cumulative logistic distribution function and  $a_j$  reflects a polynomial coefficient.

## Value

A list holding

alpha	a $T \times p \times K$ array of group-specific time-varying parameters
beta	a $T \times p \times N$ array of individual time-varying parameters
groups	a vector indicating the group memberships $(g_1,\ldots,g_N)$ , where $g_i=k$ if $i\in {\rm group}\; k$ .
У	a $NT \times 1$ vector of the dependent variable, with $\boldsymbol{y}=(y_1,\ldots,y_N)',\ y_i=(y_{i1},\ldots,y_{iT})'$ and the scalar $y_{it}$ .
X	a $NT \times p$ matrix of explanatory variables, with $\boldsymbol{X}=(x_1,\ldots,x_N)',\ x_i=(x_{i1},\ldots,x_{iT})'$ and the $p\times 1$ vector $x_{it}$ .
data	a $NT \times (p+1)$ data.frame of the outcome and the explanatory variables.

#### Author(s)

Paul Haimerl

## **Examples**

```
# Simulate a time-varying panel subject to a time trend and a group structure set.seed(1) sim <- sim_tv_DGP(N = 20, n_periods = 50, p = 1) y <- sim$y X <- sim$X
```

tv\_pagfl

Time-varying Pairwise Adaptive Group Fused Lasso

## **Description**

Estimate a time-varying panel data model with a latent group structure using the pairwise adaptive group fused lasso (*time-varying PAGFL*). The *time-varying PAGFL* jointly identifies the latent group structure and group-specific time-varying functional coefficients. The time-varying coefficients are modeled as polynomial B-splines. The function supports both static and dynamic panel data models.

# Usage

```
tv_pagfl(
  formula,
  data,
  index = NULL,
 n_periods = NULL,
  lambda,
 d = 3,
 M = floor(length(y)^(1/7) - log(p)),
 min_group_frac = 0.05,
  const_coef = NULL,
  kappa = 2,
 max_iter = 50000,
  tol_convergence = 1e-10,
  tol_group = 0.001,
  rho = 0.04 * log(N * n_periods)/sqrt(N * n_periods),
 varrho = 1,
  verbose = TRUE,
  parallel = TRUE,
)
## S3 method for class 'tvpagfl'
summary(object, ...)
## S3 method for class 'tvpagfl'
formula(x, ...)
```

```
## S3 method for class 'tvpagfl'
df.residual(object, ...)
## S3 method for class 'tvpagfl'
print(x, ...)
## S3 method for class 'tvpagfl'
coef(object, ...)
## S3 method for class 'tvpagfl'
residuals(object, ...)
## S3 method for class 'tvpagfl'
fitted(object, ...)
```

#### **Arguments**

formula a formula object describing the model to be estimated.

data a data.frame or matrix holding a panel data set. If no index variables are

provided, the panel must be balanced and ordered in the long format Y = $(Y'_1, \ldots, Y'_N)', Y_i = (Y_{i1}, \ldots, Y_{iT})'$  with  $Y_{it} = (y_{it}, x'_{it})'$ . Conversely, if data is not ordered or not balanced, data must include two index variables that de-

clare the cross-sectional unit i and the time period t of each observation.

a character vector holding two strings. The first string denotes the name of index

the index variable identifying the cross-sectional unit i and the second string represents the name of the variable declaring the time period t. The data is automatically sorted according to the variables in index, which may produce errors when the time index is a character variable. In case of a balanced panel data set that is ordered in the long format, index can be left empty if the the

number of time periods n\_periods is supplied.

the number of observed time periods T. If an index character vector is passed, n\_periods

this argument can be left empty. Default is Null.

lambda the tuning parameter determining the strength of the penalty term. Either a single

 $\lambda$  or a vector of candidate values can be passed. If a vector is supplied, a BIC-

type IC automatically selects the best fitting  $\lambda$  value.

d the polynomial degree of the B-splines. Default is 3.

the number of interior knots of the B-splines. If left unspecified, the default М

> heuristic  $M = \text{floor}((NT)^{\frac{1}{7}} - \log(p))$  is used. Note that M does not include the boundary knots and the entire sequence of knots is of length M+d+1.

min\_group\_frac the minimum group cardinality as a fraction of the total number of individuals

N. In case a group falls short of this threshold, each of its members is allocated to one of the remaining groups according to the MSE. Default is 0.05.

a character vector containing the variable names of explanatory variables that const\_coef

enter with time-constant coefficients.

kappa the a non-negative weight used to obtain the adaptive penalty weights. Default

is 2.

max\_iter the maximum number of iterations for the *ADMM* estimation algorithm. Default is  $5 * 10^4$ .

tol\_convergence

the tolerance limit for the stopping criterion of the iterative *ADMM* estimation

algorithm. Default is  $1 * 10^{-10}$ .

tol\_group the tolerance limit for within-group differences. Two individuals are assigned

to the same group if the Frobenius norm of their coefficient vector difference is

below this threshold. Default is  $1 * 10^{-3}$ .

rho the tuning parameter balancing the fitness and penalty terms in the IC that deter-

mines the penalty parameter  $\lambda$ . If left unspecified, the heuristic  $\rho=0.07\frac{\log(NT)}{\sqrt{NT}}$ 

of Mehrabani (2023, sec. 6) is used. We recommend the default.

varrho the non-negative Lagrangian ADMM penalty parameter. For the employed pe-

nalized sieve estimation *PSE*, the  $\varrho$  value is trivial. We recommend the default

1.

verbose logical. If TRUE, helpful warning messages are shown. Default is TRUE.

parallel logical. If TRUE, certain operations are parallelized across multiple cores. De-

fault is TRUE.

... ellipsis

object of class typagfl.
x of class typagfl.

#### Details

Consider the grouped time-varying panel data model

$$y_{it} = \gamma_i + \beta'_i(t/T)x_{it} + \epsilon_{it}, \quad i = 1, ..., N, \ t = 1, ..., T,$$

where  $y_{it}$  is the scalar dependent variable,  $\gamma_i$  is an individual fixed effect,  $x_{it}$  is a  $p \times 1$  vector of explanatory variables, and  $\epsilon_{it}$  is a zero mean error. The coefficient vector  $\beta_i(t/T)$  is subject to the latent group pattern

$$\beta_i\left(\frac{t}{T}\right) = \sum_{k=1}^K \alpha_k\left(\frac{t}{T}\right) \mathbf{1}\{i \in G_k\},\,$$

with 
$$\bigcup_{k=1}^K G_k = \{1, \dots, N\}$$
,  $G_k \cap G_j = \emptyset$  and  $\|\alpha_k - \alpha_j\| \neq 0$  for any  $k \neq j$ ,  $k = 1, \dots, K$ .

The time-varying coefficient functions are estimated as polynomial B-splines using the penalized sieve-technique. To this end, let B(v) denote a M+d+1 vector basis functions, where d denotes the polynomial degree and M the number of interior knots. Then,  $\beta_i(t/T)$  and  $\alpha_k(t/T)$  are approximated by forming linear combinations of the basis functions  $\beta_i(t/T) \approx \pi_i' B(t/T)$  and  $\alpha_i(t/T) \approx \xi_k' B(t/T)$ , where  $\pi_i$  and  $\xi_i$  are  $(M+d+1) \times p$  coefficient matrices.

The explanatory variables are projected onto the spline basis system, which results in the  $(M+d+1)p \times 1$  vector  $z_{it} = x_{it} \otimes B(v)$ . Subsequently, the DGP can be reformulated as

$$y_{it} = \gamma_i + z'_{it} \operatorname{vec}(\pi_i) + u_{it},$$

where  $u_{it} = \epsilon_{it} + \eta_{it}$  and  $\eta_{it}$  reflects a sieve approximation error. We refer to Su et al. (2019, sec. 2) for more details on the sieve technique.

Inspired by Su et al. (2019) and Mehrabani (2023), the time-varying PAGFL jointly estimates the functional coefficients and the group structure by minimizing the criterion

$$Q_{NT}(\boldsymbol{\pi}, \lambda) = \frac{1}{NT} \sum_{i=1}^{N} \sum_{t=1}^{T} (\tilde{y}_{it} - \tilde{z}'_{it} \text{vec}(\pi_i))^2 + \frac{\lambda}{N} \sum_{i=1}^{N-1} \sum_{j>i}^{N} \dot{\omega}_{ij} \|\pi_i - \pi_j\|$$

with respect to  $\pi = (\text{vec}(\pi_i)', \dots, \text{vec}(\pi_N)')'$ .  $\tilde{a}_{it} = a_{it} - T^{-1} \sum_{t=1}^T a_{it}$ ,  $a = \{y, z\}$  to concentrate out the individual fixed effects  $\gamma_i$ .  $\lambda$  is the penalty tuning parameter and  $\dot{w}_{ij}$  denotes adaptive penalty weights which are obtained by a preliminary non-penalized estimation.  $\|\cdot\|$  represents the Frobenius norm. The solution criterion function is minimized via the iterative alternating direction method of multipliers (*ADMM*) algorithm proposed by Mehrabani (2023, sec. 5.1).

Two individuals are assigned to the same group if  $\|\operatorname{vec}(\hat{\pi}_i - \hat{\pi}_j)\| \le \epsilon_{\text{tol}}$ , where  $\epsilon_{\text{tol}}$  is determined by tol\_group. Subsequently, the number of groups follows as the number of distinct elements in  $\hat{\pi}$ . Given an estimated group structure, it is straightforward to obtain post-Lasso estimates  $\hat{\xi}$  using group-wise least squares (see grouped\_tv\_plm).

We recommend identifying a suitable  $\lambda$  parameter by passing a logarithmically spaced grid of candidate values with a lower limit close to 0 and an upper limit that leads to a fully homogeneous panel. A BIC-type information criterion then selects the best fitting  $\lambda$  value.

In case of an unbalanced panel data set, the earliest and latest available observations per group define the start and end-points of the interval on which the group-specific time-varying coefficients are defined.

#### Value

An object of class typagfl holding

model a data.frame containing the dependent and explanatory variables as well as

cross-sectional and time indices,

coefficients let  $p^{(1)}$  denote the number of time-varying coefficients and  $p^{(2)}$  the number of

time constant parameters. A list holding (i) a  $T \times p^{(1)} \times \hat{K}$  array of the post-Lasso group-specific functional coefficients and (ii) a  $K \times p^{(2)}$  matrix of

time-constant post-Lasso estimates.

groups a list containing (i) the total number of groups  $\hat{K}$  and (ii) a vector of estimated

group memberships  $(\hat{g}_1, \dots, \hat{g}_N)$ , where  $\hat{g}_i = k$  if i is assigned to group k,

residuals a vector of residuals of the demeaned model,

fitted a vector of fitted values of the demeaned model,

args a list of additional arguments,

IC a list containing (i) the value of the IC, (ii) the employed tuning parameter  $\lambda$ ,

and (iii) the MSE,

convergence a list containing (i) a logical variable if convergence was achieved and (ii) the

number of executed ADMM algorithm iterations,

call the function call.

An object of class typagfl has print, summary, fitted, residuals, formula, df.residual and coef S3 methods.

## Author(s)

Paul Haimerl

#### References

Mehrabani, A. (2023). Estimation and identification of latent group structures in panel data. *Journal of Econometrics*, 235(2), 1464-1482. doi:10.1016/j.jeconom.2022.12.002.

Su, L., Wang, X., & Jin, S. (2019). Sieve estimation of time-varying panel data models with latent structures. *Journal of Business & Economic Statistics*, 37(2), 334-349. doi:10.1080/07350015.2017.1340299.

## **Examples**

```
# Simulate a time-varying panel with a trend and a group pattern set.seed(1)  \begin{aligned} & \text{sim} <- \text{sim}_{\text{tv}} \text{DGP}(\text{N} = 10, \text{ n_periods} = 50, intercept = TRUE, p = 1)} \\ & \text{df} <- \text{data.frame}(\text{y} = \text{c(sim$y)}) \end{aligned}  # Run the time-varying PAGFL estim <- tv_pagfl(y ~ ., data = df, n_periods = 50, lambda = 10, parallel = FALSE) summary(estim)
```

# **Index**

```
coef.gplm(grouped_plm), 2
coef.pagfl(pagfl), 9
coef.tv_gplm(grouped_tv_plm), 6
coef.tvpagfl(tv_pagfl), 19
df.residual.gplm(grouped_plm), 2
df.residual.pagfl(pagfl),9
df.residual.tv_gplm(grouped_tv_plm), 6
df.residual.tvpagfl(tv_pagfl), 19
fitted.gplm(grouped_plm), 2
fitted.pagfl(pagfl), 9
fitted.tv_gplm (grouped_tv_plm), 6
fitted.tvpagfl(tv_pagfl), 19
formula.gplm (grouped_plm), 2
formula.pagfl (pagfl), 9
formula.tv_gplm(grouped_tv_plm), 6
formula.tvpagfl(tv_pagfl), 19
grouped_plm, 2, 12
grouped_tv_plm, 6, 22
PAGFL (pagf1), 9
pagfl, 9
print.gplm (grouped_plm), 2
print.pagfl (pagfl), 9
print.tv_gplm (grouped_tv_plm), 6
print.tvpagfl (tv_pagfl), 19
residuals.gplm(grouped_plm), 2
residuals.pagfl(pagfl),9
residuals.tv_gplm(grouped_tv_plm), 6
residuals.tvpagfl(tv_pagfl), 19
sim_DGP, 14
sim_tv_DGP, 16
summary.gplm(grouped_plm), 2
summary.pagfl(pagfl),9
summary.tv_gplm(grouped_tv_plm), 6
summary.tvpagfl (tv_pagfl), 19
tv_pagfl, 19
```